

9.1. ks. 12.1.1 Napaisuudet

Välin  $[Y, aY]$  peitetodennäköisyys on

$$P(\theta \in [Y, aY]) = P(Y \leq \theta \leq aY) = P\left(\frac{1}{a} \leq \frac{Y}{\theta} \leq 1\right) = 1 - \left(\frac{1}{a}\right)^n$$

$n=5$  välin luottamustaso on siis

$$1 - \left(\frac{1}{a}\right)^5 = 0.9$$

$$\left(\frac{1}{a}\right)^5 = 0.1$$

$$\frac{1}{a} = \sqrt[5]{0.1} \Rightarrow a = \frac{1}{\sqrt[5]{0.1}} \approx 1.58$$

$$Y = \max\{2.3, 1.9, 4.5, 3.8, 2.5\} = 4.5$$

Joten 90%:n luottamusväli

$$[Y, aY] = [4.5, a \cdot 4.5] \approx [4.5, 7.13]$$

9.2. ks. Esim. 12.5

$$P(a \leq V \leq b) = 1 - \alpha$$

$$V = \frac{(n-1)S^2}{\sigma^2} \sim \chi^2(n-1)$$

Nyt  $1 - \alpha = 0.8$  eli

$$a = \chi_{0.1; n-1}^2 \quad \text{ja} \quad b = \chi_{0.9; n-1}^2 \quad n=15$$

$$= qchisq(0.1, 14) = 7.79 \quad = qchisq(0.9, 14) = 21.06$$

$$S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2 = \frac{1}{n-1} (\sum X_i^2 - n\bar{X}^2) = \frac{1}{n-1} \left( \sum X_i^2 - \frac{(\sum X_i)^2}{n} \right)$$

$$= \frac{1}{15-1} \left( 338 - \frac{63^2}{15} \right) \approx 5.24$$

$$\text{Jos } 80\% \text{ lv. } \sigma^2 \text{ lle on } \left[ \frac{(n-1)S^2}{b}, \frac{(n-1)S^2}{a} \right] \text{ eli } \left[ \frac{14 \cdot 5.24}{21.06}, \frac{14 \cdot 5.24}{7.79} \right]$$

$$= [3.48, 9.42]$$

9.3.  $\frac{2}{\theta} \sum X_i \sim \chi^2(2n)$

$$P\left(a \leq \frac{2\sum X_i}{\theta} \leq b\right) = P\left(\frac{1}{b} \leq \frac{\theta}{2\sum X_i} \leq \frac{1}{a}\right)$$

$$= P\left(\frac{2\sum X_i}{b} \leq \theta \leq \frac{2\sum X_i}{a}\right) = P\left(\frac{2n\bar{X}}{b} \leq \theta \leq \frac{2n\bar{X}}{a}\right)$$

$n=20$ ,  $\bar{X}=3.2$  a ja b  $\chi^2(2 \cdot 20)$  -jakaumasta

$$a = \chi_{0.025; 40}^2 = qchisq(0.025, 40) = 24.433$$

$$b = \chi_{0.975; 40}^2 = qchisq(0.975, 40) = 59.342$$

siis 95%:n lv.  $\theta$ :lle:

$$\left[ \frac{2 \cdot 20 \cdot 3.2}{59.342}, \frac{2 \cdot 20 \cdot 3.2}{24.433} \right] = [2.16, 5.24]$$

9.4.  $100(1-\alpha)\%$  in lv. painon muutokselle (ks. alkuun 12.2 loppu)

$$\bar{d} \pm t_{\alpha/2; n-1} \frac{S_d}{\sqrt{n}}, \text{ missä } \bar{d} = \text{havaintojen eli muutosten keskiarvo}$$

$S_d =$  muutosten hajonta

$$n = 12$$

Lastettava painon muutos

Loppupaino	73.0	73.8	75.6	80.2	76.0	80.3	74.4	79.2	73.5	76.5	75.3	74.0
Alkupaino	71.0	74.3	74.2	82.4	75.7	81.1	70.7	79.3	72.9	76.3	74.4	74.1
Erotus (d)	2.0	-0.5	1.4	-2.2	0.3	-0.8	3.7	-0.1	0.6	0.2	0.9	-0.1

$$\bar{d} = \frac{2.0 + (-0.5) + \dots + (-0.1)}{12} = \frac{5.4}{12} = 0.45$$

$$S_d^2 = \frac{1}{12-1} \sum_{i=1}^{12} (d_i - \bar{d})^2 = \frac{1}{11} \left( \sum_{i=1}^{12} d_i^2 - 12 \bar{d}^2 \right)$$

$$= \frac{1}{11} (26.7 - 12 \cdot 0.45^2) = \frac{1}{11} (26.7 - 2.43) = \frac{24.27}{11}$$

$$S_d = \sqrt{\frac{24.27}{11}} \approx 1.485$$

95% in luottamusväli:  $\alpha = 0.05$

$$0.45 \pm t_{0.025; 11} \frac{1.485}{\sqrt{12}} = (-0.49, 1.39)$$

$$t\text{-jakautumantaulukosta tai } qt(0.025, 11) = 2.201$$

9.5. (ks. luku 12.5)

havaintojen pienin arvo  $X_{(1)} = 4.9$

havaintojen suurin arvo  $X_{(12)} = 7.9$

Muotoa  $(X_{(1)}, X_{(12)})$  oleva luottamusväli on siis  $(4.9, 7.9)$

välin luottamustaso on

$$P(X_{(1)} < m < X_{(12)}) = 1 - \left(\frac{1}{2}\right)^{12-1} = 1 - \left(\frac{1}{2}\right)^{11} \approx \underline{\underline{0.9995}} \quad (\text{s. 332})$$

9.6.  $Y_i \sim N(\beta X_i, 1) \quad i=1, \dots, n$

$$l(\beta) = -\frac{n}{2} \log(2\pi) - \frac{1}{2} \sum_{i=1}^n (y_i - \beta x_i)^2 \quad (12.6.2)$$

ratkaistaan s.e.

$$l'(\beta) = -\frac{1}{2} \cdot 2 \cdot \sum_{i=1}^n (y_i - \beta x_i) \cdot (-x_i) = 0$$

$$\sum x_i y_i - \sum \beta x_i^2 = 0$$

$$\sum x_i y_i = \beta \sum x_i^2$$

merk.  $SS_x = \sum x_i^2$

$$\beta = \frac{\sum x_i y_i}{\sum x_i^2} = \frac{\sum x_i}{\sum x_i^2} y_i \Rightarrow \hat{\beta} = \frac{\sum x_i}{SS_x} y_i$$

$$\text{Var}(\hat{\beta}) = \frac{1}{SS_x^2} \sum x_i^2 \cdot \sum \text{Var}(y_i) = \frac{1}{SS_x}$$

$$\hat{\beta} \sim N\left(\beta, \frac{1}{SS_x}\right) \quad \frac{\hat{\beta} - \beta}{1/SS_x} \sim N(0, 1) \Rightarrow$$

$$\begin{aligned}
 P(-z_{\alpha/2} \leq \frac{\hat{\beta} - \beta}{\frac{1}{\sqrt{SSx}}} \leq z_{\alpha/2}) &= P(-z_{\alpha/2} \cdot \frac{1}{\sqrt{SSx}} \leq \hat{\beta} - \beta \leq z_{\alpha/2} \cdot \frac{1}{\sqrt{SSx}}) \\
 &= P(-\hat{\beta} - z_{\alpha/2} \frac{1}{\sqrt{SSx}} \leq -\beta \leq -\hat{\beta} + z_{\alpha/2} \frac{1}{\sqrt{SSx}}) \\
 &= P(\hat{\beta} + z_{\alpha/2} \frac{1}{\sqrt{SSx}} \geq \beta \geq \hat{\beta} - z_{\alpha/2} \frac{1}{\sqrt{SSx}}) \\
 &= P(\hat{\beta} - z_{\alpha/2} \frac{1}{\sqrt{SSx}} \leq \beta \leq \hat{\beta} + z_{\alpha/2} \frac{1}{\sqrt{SSx}}) = 1 - \alpha
 \end{aligned}$$

100 · (1 - α) % in lv.      α = 0.05      z<sub>α/2</sub> = z<sub>0.025</sub> = 1.96

$$\hat{\beta} \pm 1.96 \cdot \frac{1}{\sqrt{SSx}}$$

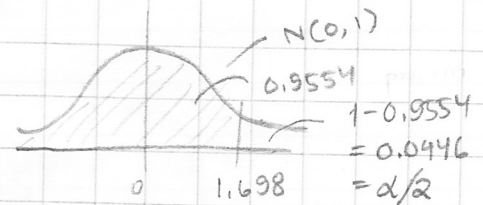
97. a) 100(1 - α) % luottamusväli π:lle  
 $\hat{\pi} \pm z_{\alpha/2} \sqrt{\frac{\hat{\pi}(1-\hat{\pi})}{n}}$  eli tässä 95% in lv. n = 800

$$0.48 \pm 1.96 \sqrt{\frac{0.48 \cdot 0.52}{800}} \quad (0.445; 0.515) \quad \cdot 0.330$$

b)  $z_{\alpha/2} \cdot \sqrt{\frac{0.48 \cdot 0.52}{800}} = 0.03$  (Luku 12.4)

∴  $z_{\alpha/2} = \frac{0.03}{\sqrt{\frac{0.48 \cdot 0.52}{800}}} = 1.698$

← maksimirike



N(0,1) taulukosta  $\Phi(1.698) \approx 0.9554$

$$\alpha/2 = 1 - 0.9554 \Rightarrow \alpha = 0.0892$$

Luottamustaso siis  $1 - \alpha = \underline{\underline{0.91}}$

98. 90% in luottamusväli  $\frac{\hat{\sigma}_1}{\hat{\sigma}_2}$  :lle

$$\frac{S_2^2 / \hat{\sigma}_2^2}{S_1^2 / \hat{\sigma}_1^2} \sim F(12, 15), \text{ missä } S_1^2 \text{ ja } S_2^2 \text{ ovat otosvariansseja}$$

$$P(a \leq \frac{S_2^2 / \hat{\sigma}_2^2}{S_1^2 / \hat{\sigma}_1^2} \leq b) = 0.90$$

$$P(a \frac{S_1^2}{S_2^2} \leq \frac{\hat{\sigma}_1^2}{\hat{\sigma}_2^2} \leq b \frac{S_1^2}{S_2^2}) = 0.90$$

$$P(\sqrt{a} \frac{S_1}{S_2} \leq \frac{\hat{\sigma}_1}{\hat{\sigma}_2} \leq \sqrt{b} \frac{S_1}{S_2})$$

$$S_1 = 0.197$$

$$S_2 = 0.318$$

$$a = F_{0.05; 12, 15} = 0.382$$

$$b = F_{0.95; 12, 15} = 2.475$$

$$\therefore \text{lv. } \left( \sqrt{0.382} \frac{0.197}{0.318}, \sqrt{2.475} \frac{0.197}{0.318} \right) \text{ eli } (0.383, 0.975)$$