

8.1) $a_1 X_1 + \dots + a_n X_n = \sum_{i=1}^n a_i X_i$ on μ in harhaton
estimaattori, jos $E(\sum_{i=1}^n a_i X_i) = \mu$

$$E(\sum_{i=1}^n a_i X_i) = \sum_{i=1}^n a_i E(X_i) = \sum_{i=1}^n a_i \mu = (\sum_{i=1}^n a_i) \mu = \mu \quad \text{käytillä } \mu \in \mathbb{R}$$

Siis on oltava

$$\sum_{i=1}^n a_i = 1$$

lisäksi asetetaan ehto
 $a_i > 0, \quad i=1, 2, \dots, n$

8.2) $\frac{\bar{X} - \mu}{\delta/\sqrt{n}} \xrightarrow{d} N(0, 1), \quad \text{kun } n \rightarrow \infty$

ts. $\sqrt{n}(\bar{X} - \mu) \xrightarrow{d} N(0, \delta^2), \quad \text{kun } n \rightarrow \infty$ (ks. Lause 11.1 ja
Esimerkki 11.8)

Nyt

$$g(\bar{X}) = \exp(\bar{X})$$

$$g(\mu) = e^\mu$$

$$g'(\mu) = e^\mu$$

$$[g'(\mu)]^2 = e^{2\mu}$$

$$\text{Var}(\bar{X}) = \frac{\delta^2}{n}$$

$$E(g(\bar{X})) = g(E(\bar{X})) = g(\mu) = e^\mu$$

$$\text{Var}(g(\bar{X})) = \text{Var}(\bar{X}) [g'(\mu)]^2 = \frac{\delta^2}{n} e^{2\mu}$$

Siis lauseen 11.1 mukaan

$\exp(\bar{X})$ noudattaa likimain normaalijakaumaa
 $N(e^\mu, \frac{\delta^2}{n} e^{2\mu})$

8.3.

$$g(\bar{X}) = \sqrt{\bar{X}}$$

$$g(\lambda) = \lambda^{\frac{1}{2}}$$

$$g'(\lambda) = \frac{1}{2} \lambda^{-\frac{1}{2}}$$

$$E(\bar{X}) = \lambda$$

$$\delta^2 = \text{Var}(\bar{X}) = \frac{\lambda}{n}$$

$$\delta^2 [g'(\lambda)]^2 = \frac{\lambda}{n} \left[\frac{1}{2} \lambda^{-\frac{1}{2}} \right]^2 = \frac{\lambda}{n} \left[\frac{1}{4} \lambda^{-1} \right] = \frac{1}{4n}$$

$$\therefore \frac{\sqrt{\bar{X}} - \sqrt{\lambda}}{\sqrt{\frac{1}{4n}}} = 2\sqrt{n}(\sqrt{\bar{X}} - \sqrt{\lambda}) \xrightarrow{d} N(0, 1)$$

ts. $\sqrt{\bar{X}} \stackrel{\text{likim.}}{\sim} N(\sqrt{\lambda}, \frac{1}{4n})$

(Lause 11.1)

8.4) Ber(θ) hav. otos $X_1=0, X_2=1, X_3=1, X_4=1$. $T = \sum_{i=1}^4 X_i$

Merk.

$$A = (X_1=0, X_2=1, X_3=1, X_4=1)$$

$$B = (T=3) = \left\{ (X_1=x_1, X_2=x_2, X_3=x_3, X_4=x_4) \mid \sum_{i=1}^4 X_i = 3 \right\}$$

ts. tapahtuma B toteutuu otoksilla

$$(0, 1, 1, 1), (1, 0, 1, 1), (1, 1, 0, 1), (1, 1, 1, 0)$$

$$P(A|B) = \frac{P(A \cap B)}{P(B)} = \frac{P(A)}{P(B)} = \frac{\theta^3(1-\theta)}{\binom{4}{3}\theta^3(1-\theta)} = \frac{1}{\binom{4}{3}} = \frac{1}{4}$$

Huom. $T = \sum_{i=1}^4 X_i \sim \text{Bin}(4, \theta)$, joten

$$P(B) = P(T=3) = \binom{4}{3}\theta^3(1-\theta)$$

$\Rightarrow P(A|B)$ ei riipu θ :sta

ja

$A \subset B$, joten $A \cap B = A$ *

Jos $B = (T=t)$ ja $t \neq 3$, niin $A \cap B = \emptyset$

ja $P(A|B) = 0 \Rightarrow$ todennäköisyys ei riipu θ :sta.

8.5) X_1, \dots, X_n otos $N(\mu, \sigma^2)$

$$f(x_i) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x_i-\mu)^2}{2\sigma^2}}$$

$$a) \sigma^2=1 \quad f(x_1, \dots, x_n; \mu) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi}} e^{-\frac{(x_i-\mu)^2}{2}} = \frac{1}{\sqrt{2\pi}^n} e^{-\frac{1}{2} \sum_{i=1}^n (x_i-\mu)^2}$$

$$= (2\pi)^{-\frac{n}{2}} \exp\left(-\frac{1}{2} \sum_{i=1}^n [(x_i-\bar{x})^2 + (\bar{x}-\mu)^2]\right)$$

$$= (2\pi)^{-\frac{n}{2}} \exp\left(-\frac{1}{2} \sum_{i=1}^n (x_i-\bar{x})^2 - \frac{n}{2} (\bar{x}-\mu)^2\right)$$

$$= \underbrace{\exp\left(-\frac{n}{2} (\bar{x}-\mu)^2\right)}_{g(\bar{x}, \mu)} \cdot \underbrace{(2\pi)^{-\frac{n}{2}} \exp\left(-\frac{1}{2} \sum_{i=1}^n (x_i-\bar{x})^2\right)}_{h(x_1, \dots, x_n)}$$

$\therefore \bar{X}$ on μ :n tyhjentävä tunnusluku

ei riipu μ :sta

$$b) \mu=0 \quad f(x_1, \dots, x_n; \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}^n} \exp\left(-\frac{\sum x_i^2}{2\sigma^2}\right)$$

$$= (2\pi\sigma^2)^{-\frac{n}{2}} \exp\left[-\frac{\sum x_i^2}{2\sigma^2}\right]$$

$$= \underbrace{(\sigma^2)^{-\frac{n}{2}} \exp\left[-\frac{\sum x_i^2}{2\sigma^2}\right]}_{g(\sum x_i^2, \sigma^2)} \cdot \underbrace{(2\pi)^{-\frac{n}{2}}}_{h(x_1, \dots, x_n)}$$

$$h(x_1, \dots, x_n) = (2\pi)^{-\frac{n}{2}}$$

ei riipu σ^2 :sta

c) $\hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n X_i^2$ on σ^2 :n harhaton estimaattori, jos $\mu=0$, koska

$$E\left(\frac{1}{n} \sum_{i=1}^n X_i^2\right) = \frac{1}{n} \sum_{i=1}^n E(X_i^2) = \frac{1}{n} \sum_{i=1}^n (\sigma^2 + \mu^2) = \sigma^2 + \mu^2 = \sigma^2, \text{ kun } \mu=0$$

$\hat{\sigma}^2$ on tyhjentävä, koska se on tyhjentävän tunnusluvun $\sum X_i^2$ funktio.

Jos $\mu \neq 0$ $S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$ on harhaton estimaattori.

$$(8.6) \quad a) \quad g(\bar{x}, \mu) = g(5, \mu) = \exp\left[-\frac{n}{2}(\bar{x} - \mu)^2\right]$$

$$L(\mu) = \text{vakio} \cdot g(5, \mu)$$

$$\ell(\mu) = -\frac{n}{2}(\bar{x} - \mu)^2$$

$$\ell'(\mu) = 2 \cdot \left(-\frac{n}{2}\right)(\bar{x} - \mu) \cdot (-1) = n(\bar{x} - \mu) = 0, \quad \text{kun } \mu = \bar{x} \\ \Rightarrow \hat{\mu} = 5$$

$$b) \quad 1. \text{ otos: } \bar{x} = \frac{3+6+4+3+9}{5} = \frac{25}{5} = 5$$

$$g(\bar{x}, \mu) = g(5, \mu) = \exp\left[-\frac{5}{2}(\bar{x} - \mu)^2\right]$$

$$h(3, 6, 4, 3, 9) = (2\pi)^{-\frac{5}{2}} \exp\left[-\frac{1}{2} \sum_{i=1}^5 (x_i - \bar{x})^2\right] \approx 0,000$$

$$2. \text{ otos: } \bar{x} = \frac{5+9+1+6+4}{5} = 5$$

$$g(\bar{x}, \mu) = g(5, \mu) = \exp\left[-\frac{5}{2}(\bar{x} - \mu)^2\right]$$

$$h(5, 9, 1, 6, 4) = (2\pi)^{-\frac{5}{2}} \exp\left[-\frac{1}{2} \sum_{i=1}^5 (x_i - \bar{x})^2\right] \approx 0,000$$

SUE $\hat{\mu} = \bar{x} = 5$ kummastakin otoksesta.

(8.7) Esim. 10.7 Lauseke (10.3.10) $\sigma^2 = 1$

$$\begin{aligned} \ell(\alpha, \beta) &= -\frac{n}{2} \log(2\pi) - \frac{1}{2} \sum (y_i - \alpha - \beta x_i)^2 \\ &= -\frac{1}{2} \sum (\alpha + \beta x_i)^2 + \underbrace{\alpha \sum y_i}_{t_1} + \underbrace{\beta \sum x_i y_i}_{t_2} - \frac{n}{2} \log(2\pi) - \frac{1}{2} \sum y_i^2 \end{aligned}$$

sits

$$L(\alpha, \beta) = \underbrace{\exp\left[-\frac{1}{2} \sum (\alpha + \beta x_i)^2 + \alpha t_1 + \beta t_2\right]}_{g(t_1, t_2, \alpha, \beta)} \cdot \underbrace{2\pi^{-\frac{n}{2}} e^{-\frac{1}{2} \sum y_i^2}}_{h(y_1, \dots, y_n)}$$

$$(8.8) \quad A(\eta) = \log(1 + e^\eta)$$

$$A'(\eta) = \frac{e^\eta}{1 + e^\eta} = \theta = E(X), \quad \text{sillä } \eta = \log \frac{\theta}{1 - \theta} \Rightarrow \theta = \frac{e^\eta}{1 + e^\eta}$$

$$\begin{aligned} A''(\eta) &= \frac{e^\eta (1 + e^\eta) - e^\eta e^\eta}{(1 + e^\eta)^2} = \frac{e^\eta (1 + e^\eta)}{(1 + e^\eta)^2} - \frac{(e^\eta)^2}{(1 + e^\eta)^2} \\ &= \theta - \theta^2 = \theta(1 - \theta) = \text{Var}(X) \end{aligned}$$